

## Example:

$$A = \{2, 4, 6\}$$

$$S = \{1, 2, 3, 4, 5, 6\}$$

$$B = \{3, 6\}$$

$$C = \{1, 2, 3, 4, 5\}$$



Uniform Prob.  
Space

$$P(A) = \frac{1}{2} \quad P(B) = \frac{1}{3} \quad P(C) = \frac{5}{6}$$

$$P(A \cap B) = \frac{1}{6} = P(A) \cdot P(B) \quad (A \text{ & } B \text{ are independent})$$

A : event that we get even number

B : event that we get a multiple of 3

$$P(A|B) = P(A) = \frac{1}{2}$$

$$P(A|C) = \frac{P(A \cap C)}{P(C)} = \frac{\frac{2}{6}}{\frac{5}{6}} = \frac{2}{5}$$

$$P(B|C) = \frac{1}{6}$$

$$P(A \cap B | C) = \frac{P(A \cap B \cap C)}{P(C)} = 0$$

$$P(A \cap B | C) \neq P(A|C) \cdot P(B|C)$$

Given that we  
rolled a multiple  
of 3, the prob.

that it's even is  
still  $\frac{1}{2}$

Conditioned on C, A & B  
are not independent  
anymore

## Independence can be tricky!

- It's possible that  $A, B, C$  are pairwise independent, but  $P(A, B, C) \neq P(A)P(B)P(C)$
- It's also possible that  $P(A, B, C) = P(A)P(B)P(C)$  but the events are not pairwise independent.

$k$ -wise independence:

$A_1, A_2, \dots, A_k$  are  $k$ -wise independent if  
for every subset  $S \subset \{1, 2, \dots, k\}$

$$P\left(\bigcap_{i \in S} A_i\right) = \prod_{i \in S} P(A_i)$$

When we say events are independent that's what we usually mean.

Coin Example: Assume we toss a coin twice

$$S = \{HH, HT, TH, TT\}$$

Assume coin is fair, and tosses are independent.

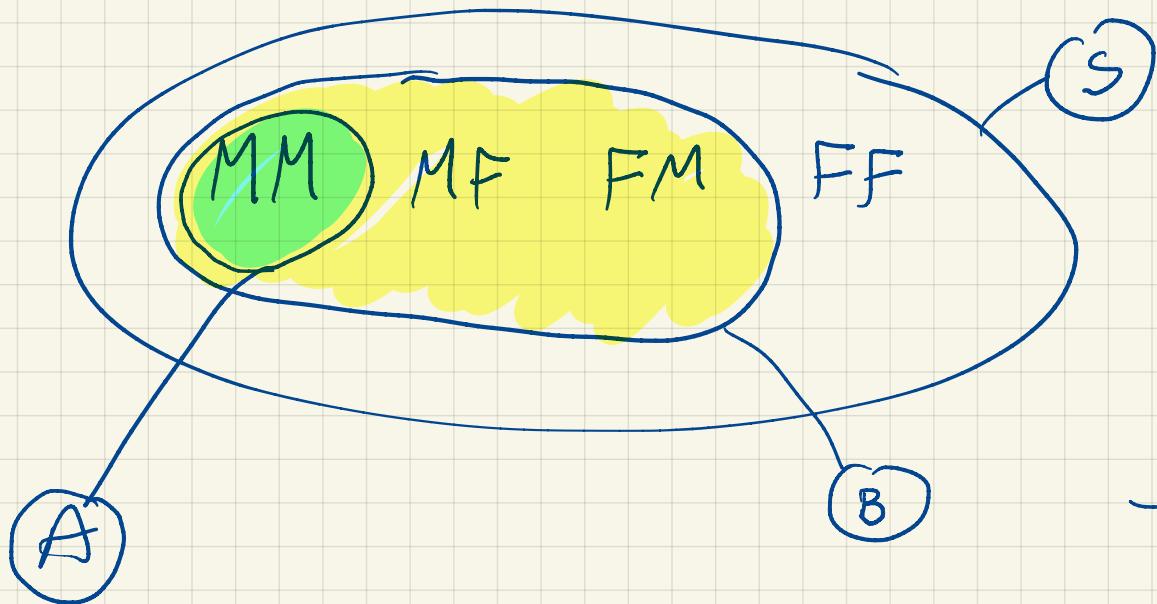
$$P(HH) = P(H)P(H) = \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4}$$

Define  $A = \{HH, HT, TH\}$      $B = \{HH, TT\}$

$$P(A) = \frac{3}{4} \quad P(B) = \frac{1}{2} \quad P(A, B) = \frac{1}{4} \quad P(A, B) \neq P(A)P(B)$$

$$P(B|A) = \frac{1}{3} \quad P(A|B) = \frac{1}{2}$$

King's Paradox: The king comes from a family of two children. What is the prob. that the king's sibling is a male?



King's Paradox

There is "implicit" conditioning in the way the question is framed.

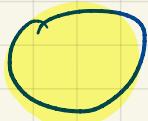
(there is at least one male!)

A: King's sibling is male

B: there is at least one male

$$P(A|B) = \frac{1/4}{3/4} = \frac{1}{3}$$

Another example:



FAIR



BIASED

$$P(H) = P(T) = \frac{1}{2}$$

$$P(H) = 1$$

- Pick one of the coins with prob.  $\frac{1}{2}$
- Toss coin 3 times.

Question: If we observe more Heads than tails what is the probability that coin is Fair.

A: Coin is Fair

$$P(A|B) = ?$$

B: we see more heads

$$B = \{ \underline{FHHH}, \underline{FHHT}, \underline{FHTH}, \underline{FTHH}, BH\bar{HH} \}$$

$$A = \{ \underline{FHHH}, \underline{FHHT}, \underline{FHTH}, \underline{FTHH}, FH\bar{TT}, FT\bar{HT}, FT\bar{TH}, FT\bar{TT} \}$$

$$P(A|B) = \frac{P(A, B)}{P(B)} = \frac{\frac{1}{16} + \frac{1}{16} + \frac{1}{16} + \frac{1}{16}}{\frac{1}{16} + \frac{1}{16} + \frac{1}{16} + \frac{1}{16} + \frac{1}{2}} = \frac{\frac{1}{4}}{\frac{1}{4} + \frac{1}{2}} = \frac{1}{3}$$

There is some subtle issues

to be clarified later regarding  
how I obtained  $\frac{1}{16}$ . (below)

$$P(FHHH) = P(F, H_1, H_2, H_3)$$

$$= \underbrace{P(F)}_{\frac{1}{2}} \underbrace{P(H_1 | F)}_{\frac{1}{2}} \underbrace{P(H_2 | H_1, F)}_{\text{Conditioned}} P(H_3 | H_1, H_2, F) \quad (\text{use chaining rule})$$

next slide

on the coin, all  
tosses are independent

↓

$$P(H_2 | H_1, F) = P(H_2 | F) = \frac{1}{2}$$

## Multiplication Rule for Conditioning:

$$P(A|B) = P(A \cap B) / P(B) \Rightarrow P(A, B) = P(B) P(A|B)$$

$$P(A, B) = P(A) \cdot P(B|A)$$

What if I have n events:

$$P(A_1, A_2, A_3, \dots, A_n) = P(A_1) P(A_2 | A_1) P(A_3 | A_1, A_2) \dots P(A_n | A_1, \dots, A_{n-1})$$

example for n=4:

$$P(A_1, A_2, A_3, A_4) = P(A_1) \cdot P(A_2 | A_1) P(A_3 | A_1, A_2) P(A_4 | A_1, A_2, A_3)$$

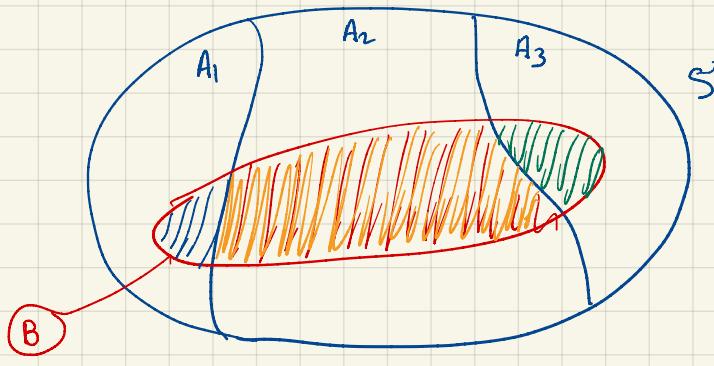
Independence when Conditioned on Coin.

$$\underbrace{P(H_1 | H_2, F)}_{=} = \frac{P(\cancel{(H_1, H_2)}, F)}{P(H_2, F)} =$$

$$\frac{P(F) P(\cancel{(H_1, H_2)} | F)}{P(H_2 | F) P(F)} = \frac{P(H_1, H_2 | F)}{P(H_2 | F)} =$$

$$= \frac{P(H_1 | F) P(H_2 | F)}{P(H_2 | F)} = \underbrace{P(H_1 | F)}_{}$$

## Law of total probability



$$B = \underbrace{(A_1 \cap B)}_{\text{exclusive}} \cup \underbrace{(A_2 \cap B)}_{\text{exclusive}} \cup \underbrace{(A_3 \cap B)}_{\text{exclusive}}$$

A's are  
1) exclusive  $A_i \cap A_j = \emptyset$

2)  $\bigcup A_i = S$  (exhaustive)

So the A's partition S

$$\begin{aligned} P(B) &= P(A_1 \cap B) + P(A_2 \cap B) + P(A_3 \cap B) \\ &= P(B|A_1)P(A_1) + P(B|A_2)P(A_2) + P(B|A_3)P(A_3) \end{aligned}$$

$$P(B) = \sum_i P(B|A_i)P(A_i)$$

example:  $P(HHH) = \underbrace{P(HHH|F)P(F)}_{\frac{1}{2}\frac{1}{2}\frac{1}{2}\cdot\frac{1}{2}} + \underbrace{P(HHH|B)P(B)}_{1\cdot1\cdot1\cdot\frac{1}{2}} = \frac{1}{16} + \frac{1}{2} = \frac{9}{16}$

## Bayes Rule

$$P(A|B) = \frac{P(A_i B)}{P(B)} = \frac{P(B|A_i) P(A_i)}{P(B)}$$

Posterior

$$P(A_i | B) = \frac{P(B|A_i) P(A_i)}{P(B) = \sum_i P(B|A_i) P(A_i)}$$

Likelihood

Prior

(Bayes Rule)

## Usual Setting:

- $A_i$ 's are not observable, but they are governed by some probabilistic law  
e.g.  $P(A_i)$  is known (prior)
- $B$  can be observed, we know how  $B$  behaves conditioned on  $A$ .

Bayes: Given what we observe, what can we say about what we don't know?