

MARKOV CHAINS.

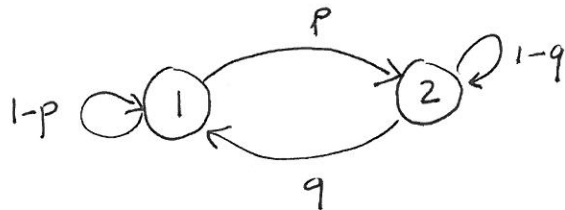
CONSIDER A SET OF STATES $S = \{1, 2, \dots, k\}$ AND ASSUME STATES EVOLVE OVER TIME IN DISCRETE TIME STEPS. LET X_n BE THE STATE AT TIME n . WE HAVE A MARKOV CHAIN IF ($m < n$):

$$P(X_n = j \mid X_0, X_1, X_2, \dots, X_m = i) = P(X_n = j \mid X_m = i)$$

FURTHER MORE WHEN $m = n-1$, THIS PROBABILITY IS GIVEN BY P_{ij} , WHERE P IS THE "TRANSITION" MATRIX, AND $\sum_{j=1}^k P_{ij} = 1$.

A SIMPLE IS GIVEN BY THE FOLLOWING MATRIX AND ALSO ILLUSTRATED BY A STATE DIAGRAM.

$$P = \begin{bmatrix} 1-p & p \\ q & 1-q \end{bmatrix}$$



AS A MOTIVATION TO STATE AN INTERESTING PROPERTY CONSIDER $P(X_2 = 2 \mid X_0 = 1)$. WE CAN REWRITE THIS AS:

$$\begin{aligned} P(X_2 = 2 \mid X_0 = 1) &= \sum_i P(X_2 = 2 \mid X_1 = i, X_0 = 1) P(X_1 = i \mid X_0 = 1) \\ &= \sum_i P(X_2 = 2 \mid X_1 = i) P(X_1 = i \mid X_0 = 1) \\ &= P(X_2 = 2 \mid X_1 = 1) P(X_1 = 1 \mid X_0 = 1) + \\ &\quad P(X_2 = 2 \mid X_1 = 2) P(X_1 = 2 \mid X_0 = 1) \\ &= p(1-p) + (1-q)p \end{aligned}$$

WE CAN OBSERVE THAT THIS IS THE ENTRY IN THE FIRST ROW AND SECOND COLUMN OF P^2

IN GENERAL,

$$P(X_n=j | X_m=l) = P_{ij}^{n-m} \quad \text{FOR } n > m.$$

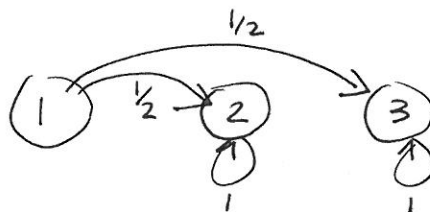
TWO PROPERTIES

THE MATRIX P IS IRREDUCIBLE IFF

$$\forall i, j \quad P(X_n=j | X_0=i) > 0 \quad \text{FOR SOME } n.$$

IN TERMS OF THE STATE DIAGRAM, THIS MEANS THERE IS A PATH FROM i TO j (NOT NECESSARILY AN IMMEDIATE TRANSITION). THE PREVIOUS MARKOV CHAIN IS IRREDUCIBLE.

THIS ONE (BELOW) IS NOT.



WHAT IS IMPORTANT ABOUT THE IRREDUCIBILITY? IT DEFINES A UNIQUE STATIONARY DISTRIBUTION OF STATES. MORE PRECISELY, IF CHAIN IS IRREDUCIBLE THEN THERE EXISTS A UNIQUE PROBABILITY VECTOR $\pi > 0$ ($\sum \pi_i = 1$) SUCH THAT

$$\pi P = \pi.$$

WHY DO WE CALL π A STATIONARY DISTRIBUTION?
 LET'S EXAMINE THIS FOR A GIVEN i .

$$[\pi_1 \ \pi_2 \ \dots \ \pi_k] \begin{bmatrix} P_{1i} \\ P_{2i} \\ \vdots \\ P_{ki} \end{bmatrix} = [\dots \ \pi_i \ \dots]$$

WE HAVE $\pi_1 P_{1i} + \pi_2 P_{2i} + \dots + \pi_k P_{ki} = \pi_i$

IF π IS A PROBABILITY DISTRIBUTION OVER STATES IN TIME n ,
 THE LEFT HAND SIDE IS THE PROBABILITY OF BEING IN STATE
 i AT TIME $n+1$. IF THIS IS EQUAL TO π_i , THEN THE PROB.
 OF STATE i HAS NOT CHANGED. THIS IS TRUE FOR ALL i .

* $P = \begin{bmatrix} 1-p & p \\ q & 1-q \end{bmatrix}$ THIS IS IRREDUCIBLE IF $p, q > 0$

$\pi = \left[\frac{q}{p+q} \quad \frac{p}{p+q} \right]$ IS STATIONARY (AND UNIQUE). VERIFY

* $P = \begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}$ THIS IS NOT IRREDUCIBLE
 ANY VECTOR OF THE FORM

$[0 \ p \ 1-p]$ IS STATIONARY.

(NOT UNIQUE)

HOW DO WE FIND THE STATIONARY DISTRIBUTION?
 WE CAN SOLVE THE SYSTEM OF EQUATIONS GIVEN BY

$$\pi P = \pi$$

USING $\sum_i \pi_i = 1$ AS AN ADDITIONAL EQUATION.

$$\begin{bmatrix} \pi_1 & \pi_2 & \dots & \pi_k \end{bmatrix} \begin{bmatrix} P \\ \mathbf{1} \end{bmatrix} = \begin{bmatrix} \pi_1 & \pi_2 & \dots & \pi_{k-1} & 1 \end{bmatrix}$$

SO LET $Q = P - I$ WITH LAST COLUMN REPLACED BY ALL ONES. THEN

$$\pi Q = [0 \ 0 \ \dots \ 0 \ 1]$$

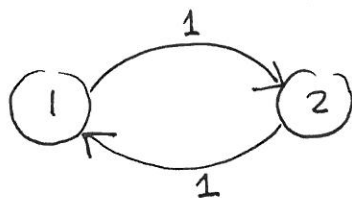
$$\pi = [0 \ 0 \ \dots \ 0 \ 1] Q^{-1}$$

A MATRIX P IS APERIODIC IF FOR EVERY i

$$\text{GCD} \{ n : P(X_n = i | X_0 = i) > 0 \} = 1$$

IN OTHER WORDS, IF WE CONSIDER THE LENGTHS OF ALL CYCLES THAT BRINGS US BACK TO i , THOSE LENGTHS HAVE NO COMMON DIVISOR (EXCEPT 1).

THIS CHAIN IS IRREDUCIBLE, BUT NOT APERIODIC



MAIN RESULT

IF A MARKOV CHAIN IS IRREDUCIBLE, THEN

$$\frac{1}{n+1} \sum_{t=0}^n f(X_t) \xrightarrow{n \rightarrow \infty} \sum_i \pi_i f(i)$$

WHERE π IS THE STATIONARY DISTRIBUTION, FOR EVERY BOUNDED FUNCTION f .

(THIS IS ANALOGOUS TO THE STRONG LAW OF LARGE NUMBERS WHERE X_0, X_1, \dots ARE INDEPENDENT. HERE THEY ARE NOT).

IN ADDITION, IF CHAIN IS APERIODIC, THEN

$$\lim_{n \rightarrow \infty} P^n = \begin{bmatrix} \pi_1 & \dots & \pi_k \\ \pi_1 & \dots & \pi_k \\ \vdots & & \vdots \\ \pi_1 & \dots & \pi_k \end{bmatrix} \quad \text{i.e.} \quad \lim_{n \rightarrow \infty} P(X_n = i | X_0) = \pi_i$$

$$\text{IN THIS CASE } \sum_i \pi_i f(i) \equiv E[f(X)]$$

REVERSIBILITY

CONSIDER THE REVERSE PROCESS GIVEN BY

$$\begin{aligned} & P(X_n = i | X_{n+1} = j, X_{n+2} = k, X_{n+3} = \dots) \\ &= \frac{P(X_n = i, X_{n+1} = j, X_{n+2} = k, X_{n+3} = \dots)}{P(X_{n+1} = j, X_{n+2} = k, X_{n+3} = \dots)} \\ &= \frac{P(X_n = i) P_{ij} P_{jk} \dots}{P(X_{n+1} = j) P_{jk} \dots} \end{aligned}$$

IF STATIONARY, THIS IS $\frac{\pi_i P_{ij}}{\pi_j}$

$$\begin{aligned} \text{So } P(X_n=i | X_{n+1}=j, X_{n+2}, X_{n+3}, \dots) \\ = P(X_n=i | X_{n+1}=j) \end{aligned}$$

So THE REVERSE PROCESS ALSO EXHIBITS THE MARKOV PROPERTY DEPENDENCE OF THE MOST RECENT STATE.

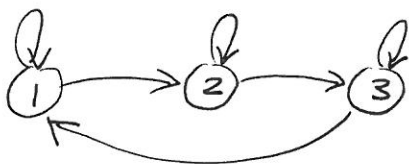
WE CALL A MARKOV CHAIN REVERSIBLE IF

$$P(X_n=i | X_{n+1}=j) = P(X_{n+1}=i | X_n=j)$$

(REVERSE TIME)

$$\text{So } \frac{\pi_i}{\pi_j} P_{ij} = P_{ji} \Rightarrow \boxed{\pi_i P_{ij} = \pi_j P_{ji}}$$

NOT ALL MARKOV CHAINS ARE REVERSIBLE. HERE'S ONE THAT IS NOT.



IF ALL TRANSITION PROBABILITIES ARE POSITIVE, THIS IS IRREDUCIBLE AND APERIODIC. BUT IF $P_{ij} > 0$, THEN $P_{ji} = 0$. SO IT CANNOT SATISFY THE REVERSIBILITY CONDITION.

WHY IS THIS IMPORTANT? IF WE FIND π SUCH THAT

$$\pi_i P_{ij} = \pi_j P_{ji} \quad \forall i, j \quad \text{THEN}$$

π IS STATIONARY.

$$\sum_i \pi_i P_{ij} = \sum_i \pi_j P_{ji} = \pi_j \sum_i P_{ji} = \pi_j$$

$$\text{So } \pi P = \pi.$$